

## Chapter 4 Summary and Review (draft: 2019/11/17-21:15:28)

### Expected value

Summary of topics and terminology:

- Expected value is the same as what we normally think of as ‘average’ or ‘mean’.
- $E(X) = \int_{-\infty}^{\infty} x f_X(x) dx$  for continuous real-valued random variable  $X$  with pdf  $f_X(x)$ .
- $E(X) = \sum_x x f_X(x)$  for discrete random variable  $X$  with pmf  $f_X(x)$ .
- Expected value of a function of RV  $X$ :  $E(u(X)) = \int_{-\infty}^{\infty} u(x) f_X(x) dx$   
or  $\sum_x u(x) f_X(x)$
- Expectation  $E()$  is a ‘linear operator’. That means it distributes over sums and differences nicely, and that we can pull out constant coefficients:
  - $E(X + Y) = E(X) + E(Y)$
  - $E(aX + b) = aE(X) + b$
  - $E(\sum_k X_k) = \sum_k E(X_k)$
  - $E(\sum_k c_k u_k(X)) = \sum_k c_k E(u_k(X))$
- Multivariate expectation:  $E(u(X, Y)) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} u(x, y) f_{X,Y}(x, y) dx dy$   
or  $\sum_x \sum_y u(x, y) f_{X,Y}(x, y)$

### **Example problems:**

1. Given pdf  $f(x) = 2x$  for  $0 < x < 1$ . Find  $E(X)$ .

*Solution:*

$$E(X) = \int_0^1 x \cdot 2x dx = \frac{2}{3} x^3 \Big|_0^1 = \frac{2}{3}.$$

2. Consider a bag with 4 red balls and 3 blue balls. Draw 5 w/o replacement. Let  $X$  be the number of blue balls drawn. Find  $E(X)$ .

*Solution:*

$$P(X = x) = \frac{\binom{4}{5-x} \binom{3}{x}}{\binom{7}{5}} \text{ for } x = 1, 2, 3.$$

$$E(X) = 1 \cdot \frac{\binom{4}{4} \binom{3}{1}}{\binom{7}{5}} + 2 \cdot \frac{\binom{4}{3} \binom{3}{2}}{\binom{7}{5}} + 3 \cdot \frac{\binom{4}{2} \binom{3}{3}}{\binom{7}{5}} = \frac{1}{7} + \frac{8}{7} + \frac{6}{7} = \frac{15}{7} \approx 2.14$$

3. Consider the joint pdf  $f(x, y) = 6e^{-2x-3y}$  for  $x > 0, y > 0$ . Calculate  $E(X + Y)$ .

*Solution:*

$$E(X+Y) = \int_0^{\infty} \int_0^{\infty} (x+y) 6e^{-2x-3y} dx dy = \int_0^{\infty} \int_0^{\infty} x 2e^{-2x} \int_0^{\infty} 3e^{-3y} dy + \int_0^{\infty} \int_0^{\infty} y 2e^{-2x} \int_0^{\infty} y 3e^{-3y} dy = \frac{1}{2} + \frac{1}{3}$$

Also notice that the joint pdf is separable and  $f(x, y) = 2e^{-2x} 3e^{-3y}$ . Thus  $X$  and  $Y$  are independent exponential random variables and  $E(X + Y) = E(X) + E(Y) = \frac{1}{2} + \frac{1}{3}$ .

## Moments

Summary of topics and terminology:

- Moments can be thought of the average value of the RV raised to a power.
- $n^{\text{th}}$  raw moment or moment about the origin:  $m_n = \mu'_n = E(X^n) = \int_{-\infty}^{\infty} x^n f_X(x) dx$  for continuous real-valued random variable  $X$  with pdf  $f_X(x)$ .
- $m_n = \mu'_n = E(X^n) = \sum_x x^n f_X(x)$  for discrete random variable  $X$  with pmf  $f_X(x)$ .
- $n^{\text{th}}$  central moment or moment about the mean:  $\mu_n = E((X - \mu)^n) = \int_{-\infty}^{\infty} (x - \mu)^n f_X(x) dx$  for continuous and  $\mu_n = E((X - \mu)^n) = \sum_x (x - \mu)^n f_X(x)$  for discrete
- Multivariate moments:  $\mu'_{n,k} = E(X^n Y^k) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x^n y^k f_{X,Y}(x, y) dx dy$   
or  $\sum_x \sum_y x^n y^k f_{X,Y}(x, y)$
- Variance is the  $2^{\text{nd}}$  central moment:  $\text{Var}(X) = \sigma_X^2 = E[(X - \mu)^2] = E(X^2) - \mu_X^2$ .
- Moment generating function:  $M_X(t) = E(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f_X(x) dx$  or  $\sum_x e^{tx} f_X(x)$
- MGF is a power series with moments as coefficients:  
 $M_X(t) = 1 + E(X)t + \frac{1}{2}E(X^2)t^2 + \frac{1}{3!}E(X^3)t^3 + \dots$
- Moments can be calculated from MGF by taking derivatives and evaluating at  $t = 0$ :  
 $E(X^n) = M_X^{(n)}(0) = \left. \frac{d^n}{dt^n} M_X(t) \right|_{t=0}$ .
- Chebyshev's theorem:  $P(|X - \mu| > k\sigma) \leq \frac{1}{k^2}$ .

### **Example problems:**

1. Given pdf  $f(x) = 2x$  for  $0 < x < 1$ . Find  $\mu, \mu'_2, \mu_2, \sigma^2$ .

Solution:

$$\mu = E(X) = \int_0^1 x \cdot 2x dx = \left. \frac{2}{3}x^3 \right|_0^1 = \frac{2}{3}.$$

$$\mu'_2 = E(X^2) = \int_0^1 x^2 \cdot 2x dx = \left. \frac{2}{4}x^4 \right|_0^1 = \frac{1}{2}.$$

$$\sigma^2 = \mu_2 = E(X^2) - [E(X)]^2 = \mu'_2 - \mu^2 = \frac{1}{2} - (2/3)^2 = \frac{1}{18}.$$

So this random variable has mean  $\frac{2}{3}$  and variance  $\text{Var}(X) = \frac{1}{18}$ .

2. Find the moment generating function for  $X$  with pdf  $f(x) = 2x$  on  $[0, 1]$ .

Solution:

$$M(t) = E(e^{tX}) = \int_0^1 e^{tx} 2x dx = 2x \frac{1}{t} e^{tx} \Big|_0^1 - 2 \int_0^1 \frac{1}{t} e^{tx} dx = 2 \frac{1}{t} e^t - 2 \frac{1}{t^2} e^{tx} \Big|_0^1 = 2 \left( \frac{1}{t} e^t - \frac{1}{t^2} e^t + \frac{1}{t^2} \right)$$

3. Given moment generating function  $M_X(t) = \frac{1}{1-t} e^{t^2}$  find  $E(X)$  and  $E(X^2)$ .

Solution:

$$M'(t) = (1-t)^{-2} e^{t^2} + (1-t)^{-1} 2te^{t^2} \text{ thus } M'(0) = 1 = E(X).$$

$$M''(t) = 2(1-t)^{-3} e^{t^2} + (1-t)^{-2} 2te^{t^2} + (1-t)^{-2} 2te^{t^2} + (1-t)^{-1} 2e^{t^2} + (1-t)^{-1} 4t^2 e^{t^2} \text{ thus } M''(0) = 4 = E(X^2).$$

## Linear combinations and covariance

Summary of topics and terminology:

- Linear combination of RVs  $X_1, X_2, \dots$  is  $\sum_{i=1}^{\infty} c_i X_i$
- $\sigma_{XY} = \text{Cov}(X, Y) = E((X - \mu_X)(Y - \mu_Y)) = E(XY) - \mu_X \mu_Y$  and can be positive or negative or zero.
- RVs  $X$  and  $Y$  independent implies that they have zero covariance, but zero covariance does not imply independence.
- $Y_1 = \sum_i a_i X_i$  and  $Y_2 = \sum_i b_i X_i$ , then:
  - $\text{Var}(Y_1) = \sum_i a_i^2 \text{Var}(X_i) + 2 \sum_{i < j} a_i a_j \text{Cov}(X_i, X_j)$
  - $\text{Cov}(Y_1, Y_2) = \sum_i a_i b_i \text{Var}(X_i) + \sum_{i < j} (a_i b_j + a_j b_i) \text{Cov}(X_i, X_j)$

### **Example problems:**

1. Let  $Y_1 = 3X_1 - X_2$  and  $Y_2 = X_1 + 5X_2$ . Given that  $X_1$  and  $X_2$  have variances, 1 and 2, respectively, and covariance -1, find the covariance of  $Y_1$  and  $Y_2$ .

Solution:

$$\begin{aligned}\text{Cov}(Y_1, Y_2) &= 3 \cdot 1 \cdot \text{Var}(X_1) + (-1) \cdot 5 \cdot \text{Var}(X_2) + (3 \cdot 5 + (-1) \cdot 1) \text{Cov}(X_1, X_2) \\ &= 3 \cdot 1 - 5 \cdot 2 + 14 \cdot (-1) = -21\end{aligned}$$

2. Given joint pdf  $f(x, y) = 2$  on  $0 < x < y < 1$ , find  $\text{Cov}(X, Y)$ .

Solution:

$$\mu_X = \int_{-\infty}^{\infty} x f(x, y) dx dy = \int_0^1 \int_0^y 2x dx dy = \frac{1}{3}$$

$$\mu_Y = \int_{-\infty}^{\infty} y f(x, y) dx dy = \int_0^1 \int_0^y 2y dx dy = \frac{2}{3}$$

$$E(XY) = \int_{-\infty}^{\infty} xy f(x, y) dx dy = \int_0^1 \int_0^y 2xy dx dy = \frac{1}{4}$$

$$\sigma_{XY} = \text{Cov}(X, Y) = E(XY) - \mu_X \mu_Y = \frac{1}{4} - \frac{1}{3} \cdot \frac{2}{3} = \frac{1}{36}$$

3. If  $\text{Cov}(X, Y) = 0$ , are  $X$  and  $Y$  independent?

Solution:

No! Independent RVs do have zero covariance, but zero covariance does not imply independence!

## Conditional expectation

Summary of topics and terminology:

- $E(X | X \in A) = \frac{\int_A x f_X(x) dx}{P(X \in A)}$
- $E(X | Y \in A) = \frac{\int_{-\infty}^{\infty} \int_A x f(x, y) dy dx}{P(Y \in A)}$
- $E(X | Y = y) = \int_{-\infty}^{\infty} x f_{X|Y=y}(x|y) dx = \int_{-\infty}^{\infty} x \frac{f(x, y)}{f_Y(y)} dx$  for  $y$  a \*fixed\* value
- The above formulas are given for continuous RVs, but the formulas for discrete RVs would be the same but with summations instead of integrals.
- We can take the expectation of functions of RVs this way as well, e.g.  
 $E(u(X) | X \in A) = \frac{\int_A u(x) f_X(x) dx}{P(X \in A)}$
- $E(u(X, Y) | X \in A, Y \in B) = \frac{\int_B \int_A u(x, y) f(x, y) dx dy}{P(X \in A, Y \in B)}$
- Conditional mean:  $\mu_{X|y} = E(X | Y = y)$
- Conditional variance:  $\sigma_{X|y}^2 = E(X^2 | Y = y) - \mu_{X|y}^2$

### **Example problems:**

1. Consider pmf  $f(x) = x/6$  for  $x = 1, 2, 3$ . Calculate  $E(X | X \neq 3)$ .

*Solution:*

$$E(X | X \neq 3) = \frac{1 \cdot \frac{1}{6} + 2 \cdot \frac{2}{6}}{\frac{1}{6} + \frac{2}{6}} = \frac{5}{3}$$

2. Given pdf  $f(x) = \frac{2}{x^3}$  for  $x > 1$ , find  $E(X | X > t)$ .

*Solution:*

$$E(X | X > t) = \frac{\int_t^{\infty} x \frac{2}{x^3} dx}{\int_t^{\infty} \frac{2}{x^3} dx} = \frac{-2x^{-2}|_t^{\infty}}{-x^{-2}|_t^{\infty}} = \frac{2/t}{1/t^2} = 2t$$

This means that, given that we know  $X$  is beyond some threshold value, we expect it to be twice that threshold.

3. Consider joint pdf  $f(x, y) = \frac{x}{y^2} + \frac{1}{2}e^{1-y}$  in  $[0, 1] \times [1, \infty)$ . Calculate the expected value of  $X$  given that  $Y$  is less than 2. Also calculate the conditional expectation of  $X$  given that  $Y$  is exactly 2.

*Solution:*

$$\begin{aligned} E(X | Y < 2) &= \frac{\int_0^1 \int_1^2 x \left( \frac{x}{y^2} + \frac{1}{2}e^{1-y} \right) dy dx}{\int_0^1 \int_1^2 \left( \frac{x}{y^2} + \frac{1}{2}e^{1-y} \right) dy dx} \\ &= \frac{\frac{1}{6} + \frac{e-1}{4e}}{\frac{1}{4} + \frac{e-1}{2e}} = \frac{4e + 6(e-1)}{6e + 12(e-1)} = \frac{5e-3}{9e-6} \approx 0.5736 \end{aligned}$$

The marginal for  $Y$  is  $f_Y(y) = \frac{1}{2y^2} + \frac{1}{2}e^{1-y}$ .

$$\begin{aligned} E(X | Y = 2) &= \int_0^1 x \frac{f(x, 2)}{f_Y(2)} dx = \int_0^1 x \frac{\frac{x}{4} + \frac{1}{2}e^{-1}}{\frac{1}{8} + \frac{1}{2}e^{-1}} dx = \frac{1}{1 + 4e^{-1}} \int_0^1 2x^2 + 4xe^{-1} dx \\ &= \frac{1}{1 + 4e^{-1}} \left[ \frac{2}{3}x^3 + 2x^2e^{-1} \right]_0^1 = \frac{\frac{2}{3} + 2e^{-1}}{1 + 4e^{-1}} \approx 0.5674 \end{aligned}$$