

Regression: ch. 14

X & Y are jointly distributed w/
w/ $f_{X,Y}(x,y)$ joint pdf.

Def. 14.1

$$\mu_{Y|X=x} = E(Y | X=x)$$

is the regression eq.
for Y given X .

$\mu_{Y|X}$ is a funct. of x

e.g. $\mu_{Y|X} = \alpha + \beta x \quad \alpha, \beta \in \mathbb{R}$

$\mu_{Y|X} = c \cdot e^{rx} \quad c, r \in \mathbb{R}$

Linear

See Math 421 notes on conditional expectation & distributions.

§4.8

e.g. $f_{Y|X=x}(y|x) = \frac{f(x,y)}{f_X(x)}$

§3.7

is the conditional pdf of Y given $X=x$

ex] $f(x,y) = 2x + 3y^2$ on $[0,1]^2$

In general:

$$\mu_{Y|X} = E(Y|X=x) = \int_0^1 y \cdot f_{Y|X=x}(y) dy$$

$$= \frac{\int_0^1 y \cdot f(x,y) dy}{\int_0^1 f(x,y) dy}$$

$$= \frac{\int_0^1 y \cdot (2x + 3y^2) dy}{\int_0^1 (2x + 3y^2) dy}$$

$$= \frac{[xy^2 + \frac{3}{4}y^4]_0^1}{[2xy + y^3]_0^1} = \frac{x + \frac{3}{4}}{2x + 1}$$

Thus the regression eq. for Y given $X=x$

$$\mu_{Y|X} = \frac{x + \frac{3}{4}}{2x + 1}$$

Non-linear
funct. of x

Note: If we had actual (X,Y) data, the regression function is not necessarily a "good fit". We'd need to investigate the conditional variance

$$\text{Var}(Y|X=x) = E(Y^2|x) - [E(Y|x)]^2$$

Linear Regression - §14.2

$$M_{Y|X=x} = \alpha + \beta x$$

Let μ_1, μ_2 be means &
 σ_1^2, σ_2^2 variances of
 X & Y

Remember: X & Y are
jointly distributed

§4.6 → Recall Covariance

$$\begin{aligned} \text{Cov}(X, Y) &= E((X - \mu_1)(Y - \mu_2)) \\ &= \sigma_{12} \end{aligned}$$

Correlation Coef: "rho" → $\rho = \frac{\sigma_{12}}{\sigma_1 \cdot \sigma_2}$

Thm 14.1 If $M_{Y|X}$ is linear in x , then:

$$M_{Y|X} = \mu_2 + \rho \cdot \frac{\sigma_2}{\sigma_1} (x - \mu_1)$$

Proof of Thm 14.1

$$M_{Y|X} = \alpha + \beta X = \frac{\int_{-\infty}^{\infty} y \cdot f(x, y) dy}{f_X(x)}$$

so:

$$\alpha \cdot f_X(x) + \beta \cdot x \cdot f_X(x) = \int_{-\infty}^{\infty} y \cdot f(x, y) dy$$

★ integrate w.r.t. x

$$\alpha + \beta \cdot E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y \cdot f(x, y) dy dx$$

$$= \int_{-\infty}^{\infty} y \int_{-\infty}^{\infty} f(x, y) dx dy$$

$$= \int_{-\infty}^{\infty} y \cdot f_Y(y) dy$$

$$= E(Y)$$

Thus $\alpha + \beta \cdot M_1 = M_2$
mult. ★ by x before integrating:

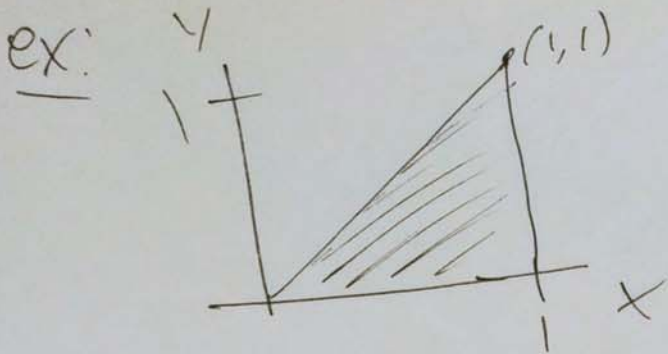
$$\underbrace{\alpha \cdot E(X)}_{M_1} + \beta \cdot \underbrace{E(X^2)}_{\sigma_1^2 + M_1^2} = \underbrace{E(XY)}_{\sigma_{12} + M_1 M_2}$$

Then have 2 eq. for α and β ★

$$\text{get: } \alpha = M_2 - \frac{\sigma_{12}}{\sigma_1^2} M_1$$

$$\beta = \frac{\sigma_{12}}{\sigma_1^2}$$

★ a bit of algebra gives Thm 14.1



joint pdf:

$$f(x,y) = z \text{ on this triangle}$$

Marginals: $f_x(x) = \int_0^x z dy = 2x \quad x \in (0,1)$

$$f_y(y) = \int_y^1 z dx = z(1-y) \quad y \in (0,1)$$

So: Regr. y given x is:

$$M_{Y|X} = \frac{\int_0^x y \cdot z dy}{2x} = \frac{[y^2]_0^x}{2x} = \frac{x^2}{2x} = \boxed{\frac{x}{2}}$$

thus $\alpha = 0, \beta = \frac{1}{2}$

but must also have $M_{Y|X} = M_2 + \rho \frac{\sigma_2}{\sigma_1} (x - M_1)$

$$M_1 = E(X) = \int_0^1 x \cdot 2x dx = \frac{2}{3}$$

$$M_2 = E(Y) = \int_0^1 y \cdot 2(1-y) dy = \frac{1}{3}$$

$$E(X^2) = \int_0^1 x^2 \cdot 2x dx = \frac{1}{2}$$

$$E(Y^2) = \int_0^1 y^2 \cdot 2(1-y) dy = \frac{1}{6}$$

$$E(XY) = \int_0^1 \int_0^x xy \cdot 2 dy dx = \int_0^1 x \cdot [y^2]_0^x dx = \int_0^1 x^3 dx = \frac{1}{4}$$

Thus $\sigma_1^2 = \sigma_2^2 = \frac{1}{18}$

$$\rho = \frac{\sigma_{12}}{\sigma_1 \sigma_2} = \frac{\frac{1}{4} - \frac{2}{3} \cdot \frac{1}{3}}{\frac{1}{18}} = \left(\frac{1}{2}\right)$$

So: $M_{Y|X} = \frac{1}{3} + \frac{1}{2} \cdot \frac{\sqrt{\frac{1}{18}}}{\sqrt{\frac{1}{18}}} (x - \frac{2}{3}) = \frac{1}{2}x$ ✓

both eq. same.